

# Military Simulation Analytics: Towards a Sensitivity Analysis when Conducting Analysis of Alternatives

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### Improving Force Design Force Structure Review

Force Structure Review (FSR) process:

Senior leadership applying military judgement over force options through seminar wargaming



Complexity of modern ops → difficult to rely on intuition for Force Design

- Many factors affect modern ops
- Difficulty in assessing impact of new capability (yet to be developed)
- Future wars fought differently to past

SR2 will deliver a sim capability for exploring & developing complex whole-of-force operating concepts



Operating Concepts for Exploration

- Force Level Electronic Warfare
- Maritime Force Defence
- Space Concepts
- Cooperative Engagement
- Information Age Combat Model
- Cyber Warfare
- **New Operating Concepts**

#### Involved Methods & Fields of Study

### Future Operating

- 1. FLEW: Force Level EW
  2. CEC: Cooperative
- Engagement Capability
- 3. Space Concepts
- 4. Maritime Force Defence
- 5. Cyber Warfare
- 6. New Concept 2

#### M&S

- Develop novel modelling strategies to represent abstract concepts in HPCsim
- Resolving computational intractability in large scale simulation (many factors)



DoE

### Analysis & Visualisation

- Develop new analysis strategies for high dim problem spaces (big data)
  - Many response vars.
  - Many design points
    Many iterations @ a
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### **Sensitivity Analysis and Analysis of Alternatives**

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#### Sensitivity Analysis:

- Global sensitivity Stochastic Kriging or Gaussian Process (optimisation)
- Local sensitivity low-order polynomials (main effects/two-way interactions)
- Sub-system marginal contribution to operational effectiveness
- Combat multipliers (combined arms combat)
- Robustness of point-scenario insights to uncertain parameters



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#### Analysis of Alternatives:

- Discrete set of competing systems (e.g. tender evaluation)
- All pairwise comparisons generally produces only partial order
- Selection of the best or subset containing the best
- Ranking analysis score-based or partition-based





### **Linear Regression**

With  $m_i$  replications at each design point  $\mathbf{x}_i$  fit using **OLS criterion**. The sum of squared residuals is:

$$SSR = \sum_{i=1}^{n} \sum_{r=1}^{m_i} (\hat{y}_i - w_{ir})^2 \text{ where } \hat{y}_i = \sum_{j=1}^{q} x_{ij} \hat{\beta}_j, i = 1, \dots, n$$

$$= \sum_{i=1}^{n} \sum_{r=1}^{m_i} \left[ (x_{ik} \hat{\beta}_k)^2 + 2x_{ik} \hat{\beta}_k \sum_{j=1; \neq k}^{q} x_{ij} \hat{\beta}_j + \left( \sum_{j=1; \neq k}^{q} x_{ij} \hat{\beta}_j \right)^2 - 2w_{ir} \sum_{j=1}^{q} x_{ij} \hat{\beta}_j + (w_{ir})^2 \right].$$

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Differentiating the SSR with respect to the k—th regression parameter gives:

$$\frac{\partial SSR}{\partial \hat{\beta}_k} = 2 \sum_{i=1}^n x_{ik} m_i \sum_{i=1}^q x_{ij} \hat{\beta}_j - 2 \sum_{i=1}^n x_{ik} m_i \overline{w}_i, \quad k = 1, \dots, q \quad \text{where} \quad \overline{w}_i = \sum_{r=1}^{m_i} w_{ir} / m_i.$$

**Normal equations** for the OLS estimator: $X'MX\hat{\beta}^{OLS} = X'M\overline{\mathbf{w}}$ .

#### **Parameter Confidence Intervals**

Since 
$$\hat{\beta}_i^{OLS} = \sum_{i=1}^n L_{ji} \overline{w}_i$$
 where  $L = (X'MX)^{-1} X'M$  and treating  $\overline{w}_i$  as **random variables**:

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$$= \sum_{i=1}^{n} \sum_{i'=1}^{n} L_{ji}L_{ji'} \times \sum_{r=1}^{\min(m_{i}, m_{i'})} cov(W_{ir}, W_{i'r})/(m_{i}m_{i'})$$

$$= \sum_{i=1}^{n} \sum_{i'=1}^{n} L_{ji}L_{ji'}\sigma_{ii'}/\max(m_{i}, m_{i'}) \text{ where } \sigma_{ii'} = cov(W_{i}, W_{i'}).$$

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This generalises and simplifies Kleijnen (2015) who treated  $m_i = m$  and  $m_i \neq m$  separately.

The covariance matrix simplifies  $\Sigma_w = \sigma^2(w) \mathbf{I}$  and either:

• 
$$\sigma^2(W) \approx s^2(w) = \sum_{i=1}^n (m_i - 1) s^2(w_i) / (N - n)$$
 pooling  $n$  sample variance estimators, or

• 
$$\sigma^2(W) \approx MSR = \sum_{i=1}^n \sum_{r=1}^{m_i} (\hat{y}_i - w_{ir})^2 / (N-q)$$
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However Kleijnen (2015) used:

"MSR" 
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Kleijnen's (2015) lack-of-fit F-statistics are incorrect. The **correct**, **general**, **expression** is:

$$F_{N-q,N-n} = \frac{MSR}{s^2(w)} = \frac{\sum_{i=1}^n \sum_{r=1}^{m_i} (w_{ir} - \hat{y}_i)^2 / (N-q)}{\sum_{i=1}^n \sum_{r=1}^{m_i} (w_{ir} - \overline{w}_i)^2 / (N-n)}.$$

With CRN as a Variance Reduction Technique  $m_i = m$  and  $\Sigma_w$  approximated by sample covariance so:

$$var(\hat{\beta}_{j}^{OLS}) \approx \sum_{i=1}^{n} \sum_{i'=1}^{n} L_{ji} L_{ji'} \sum_{r=1}^{m} (w_{ir} - \overline{w}_{i}) (w_{i'r} - \overline{w}_{i'}) / [m(m-1)].$$

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Kleijnen (2015) proposed alternative method inspired by classical text Law (2007).

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#### The two methods are identical, not alternatives.

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An approach suggested by Kleijnen (2015) is to use Jack-Knifing (Tukey, 1958):

$$J_{j;r}=m\hat{\beta}_j-(m-1)\hat{\beta}_{j;-r} \text{ where } \hat{\beta}_{j;-r}=\sum_{i=1}^n L_{ji}\overline{w}_{i;-r} \text{ and } \overline{w}_{i;-r}=\sum_{r'=1:r\neq r}^m \frac{w_{ir'}}{m-1}.$$

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So the r-th jackknifed pseuduovalue  $J_{j;r}$  is identical to OLS estimator based on r-th replication  $\mathbf{w}_r$ .

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- then Law (2007) approach is in fact identical to both OLS and Jack-Knifing in Kleijnen (2015).
- Confidence intervals for the regression coefficients can be calculated from:

$$\hat{\beta}_{j} = \sum_{i=1}^{n} L_{ji}\overline{w}_{i}, \quad j = 1, \dots, q$$

$$var(\hat{\beta}_{j}) \approx \sum_{i=1}^{n} \sum_{j'=1}^{n} L_{ji}L_{ji'} \sum_{r=1}^{m} (w_{ir} - \overline{w}_{i})(w_{i'r} - \overline{w}_{i'})/[m(m-1)], \quad j = 1, \dots, q$$

$$L = (X'X)^{-1}X' \quad \text{and} \quad \overline{w}_{i} = \sum_{r=1}^{m} w_{ir}/m, \quad i = 1, \dots, n.$$

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- Individual alternative simulation output might be sensitive, but comparatively insensitive.
- Logistic regression (two alternative case)  $P(Z = 1|\mathbf{x}) = (1 + \exp[-\beta^T \mathbf{x}])^{-1}$ ?
- Generate sample data  $z_i=1$  if reject  $H_0: \mu_{1i}=\mu_{0i}$  in favour of  $H_1: \mu_{1i}>\mu_{0i}$ .
- $odds(\mathbf{x}) = P(Z = 1|\mathbf{x})/[1 P(Z = 1|\mathbf{x})]$  and  $\exp(\hat{\beta}_i) = odds(\mathbf{x} + \mathbf{e}_i)/odds(\mathbf{x})$ .

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#### Simple counterexample:

- If  $P(Z=1|\mathbf{x})=1/3$  and  $\hat{\beta}_i=1 \to odds(\mathbf{x})=1$ : 2 and  $odds(\mathbf{x}+\mathbf{e}_i)=2.72$ : 2 = 1.36: 1.
- Preference (decision) has changed from alternative 0 to alternative 1.
- If, however,  $P(Z=1|\mathbf{x})=2/3$  and  $\hat{\beta}_j=1$ . Then  $odds(\mathbf{x})=2$ : 1 and  $odds(\mathbf{x}+\mathbf{e}_j)=5.44$ : 1.
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Unlikely that logistic regression (or **multinomial regression** for the general K > 2 case) will work.

### **Ranking Sensitivity Measure**

Score-based method of Villacorta & Sáez (2015):

$$s_{ki} = \sum_{j=k+1}^{K} z_{kji} \quad \text{where} \quad z_{kji} = \begin{cases} +1, & \text{if accept } H_1: \mu_{ki} > \mu_{ji} \\ 0, & \text{if accept } H_0: \mu_{ki} = \mu_{ji} \\ -1, & \text{if accept } H_1: \mu_{ki} < \mu_{ji} \end{cases}$$

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How to measure **similarity/distance** between  $\mathbf{s}_i$  and  $\mathbf{s}_{i'}$ ?

- Convert to ranks:  $\sigma_i(k) = rank(s_{ki}, \mathbf{s}_i)$  so  $\sigma_i(\cdot)$  is a permutation of  $\{1, \ldots, K\}$ .
- Weighted Spearman's Footrule:  $\delta_{ii'}^F = \sum_{k=1}^K d_{ii'k} p_{ii'k} |\sigma_i(k) \sigma_{i'}(k)|$ . (Dolgun *et al.*, 2018).
- Inter-rater disagreement:  $d_{ii'k} = \left(\frac{|s_{ki} s_{ki'}|}{2(K-1)}\right)^p$  (Gwet, 2014).
- Head or Tail Position:  $p_{ij'k} = \left(\frac{\sigma_i(k) + \sigma_{j'}(k)}{2}\right)^{-1/K}$  (Kumar & Vassilvitskii, 2000).

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Score-based method of Villacorta & Sáez (2015):

$$s_{ki} = \sum_{j=k+1}^{K} z_{kji}$$
 where  $z_{kji} = \begin{cases} +1, & \text{if accept } H_1: \mu_{ki} > \mu_{ji} \\ 0, & \text{if accept } H_0: \mu_{ki} = \mu_{ji} \\ -1, & \text{if accept } H_1: \mu_{ki} < \mu_{ji} \end{cases}$ 

How to measure **similarity/distance** between  $\mathbf{s}_i$  and  $\mathbf{s}_{i'}$ ?

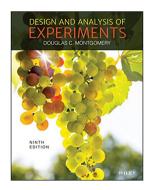
- Convert to ranks:  $\sigma_i(k) = rank(s_{ki}, s_i)$  so  $\sigma_i(\cdot)$  is a permutation of  $\{1, \ldots, K\}$ .
- Weighted Spearman's Footrule:  $\delta_{ii'}^F = \sum_{k=1}^K d_{ii'k} p_{ii'k} |\sigma_i(k) \sigma_{i'}(k)|$ . (Dolgun *et al.*, 2018).
- Inter-rater disagreement:  $d_{ii'k} = \left(\frac{|s_{ki} s_{ki'}|}{2(K-1)}\right)^p$  (Gwet, 2014).
- Head or Tail Position:  $p_{ii'k} = \left(\frac{\sigma_i(k) + \sigma_{i'}(k)}{2}\right)^{-1/K}$  (Kumar & Vassilvitskii, 2000).

This provides a scalar measure of **sensitivity of ranking vector** between two design points,  $\mathbf{x}_i$  and  $\mathbf{x}_{i'}$ .

How do we use that to isolate the sensitivity of ranking vector to an individual parameter x<sub>i</sub>?

	<b>X</b> .1	<b>X</b> .2		$\mathbf{x}_{.j}$	• • •	$\mathbf{x}_{.K}$	У
<b>x</b> <sub>1.</sub>		• • •		• • •	• • •	• • •	
		• • •	• • •	• • •	• • •	• • •	
$\mathbf{x}_{i_j+.}$	<i>X<sub>ij</sub></i> 1	$x_{i_j2}$	• • •	+1	• • •	$x_{i_jK}$	$y_{i_j+}$
		• • •	• • •	• • •	• • •	• • •	
$\mathbf{x}_{i_j}$	<i>Xi<sub>j</sub></i> 1	$x_{i_j2}$	• • •	-1	• • •	$x_{i_jK}$	y <sub>ij</sub> _
		• • •	• • •		• • •		
$\mathbf{x}_{2^{K}}$ .		• • •	• • •	• • •	• • •	• • •	





	<b>x</b> .1	<b>X</b> .2	• • •	$\mathbf{x}_{.j}$	• • •	$\mathbf{x}_{.K}$	У
$\mathbf{x}_{1.}$				• • •	• • •		
• • •		• • •		• • •	• • •	• • •	
$\mathbf{x}_{i_j+.}$	<i>X</i> <sub><i>i</i><sub>j</sub> 1</sub>	$x_{i_j2}$	• • •	+1	• • •	$x_{i_jK}$	$y_{i_j+}$
• • •	• • • •	• • •	• • •	• • •	• • •	• • •	
$\mathbf{x}_{i_j}$ .	<i>X</i> <sub><i>i</i><sub>j</sub> 1</sub>	$x_{i_j2}$	• • •	-1	• • •	$x_{i_jK}$	y <sub>ij</sub> _
• • •					• • •		
$\mathbf{X}_{2}^{\kappa}$ .							

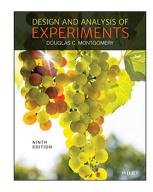
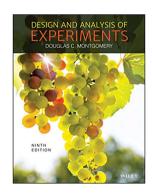


Table: Full Factorial Design for Univariate Response

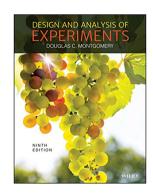
• For **orthogonal** designs  $\hat{\beta} = (X'X)^{-1}X'$ **y** becomes  $\hat{\beta}_j = \mathbf{x}'_{.j}\mathbf{y}/2^K$ .

	<b>x</b> .1	<b>X</b> .2	• • •	$\mathbf{x}_{.j}$	• • •	$\mathbf{x}_{.K}$	У
$\mathbf{x}_{1.}$				• • •	• • •		
• • •							
$\mathbf{x}_{i_j+.}$	<i>x</i> <sub>ij</sub> 1	<i>x</i> <sub>ij2</sub>		+1	• • •	$x_{i_jK}$	<i>y</i> <sub>ij</sub> +
• • •		• • •	• • •		• • •	• • •	
$\mathbf{x}_{i_j}$ .	<i>X<sub>ij</sub></i> 1	<i>X</i> <sub><i>i</i><sub>j</sub>2</sub>		-1	• • •	$x_{i_jK}$	Уij—
• • •		• • •	• • •	• • •	• • •	• • •	
$\mathbf{X}_{2}^{\kappa}$ .							



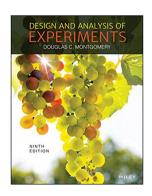
- For **orthogonal** designs  $\hat{\beta} = (X'X)^{-1}X'$ **y** becomes  $\hat{\beta}_i = \mathbf{x}'_i\mathbf{y}/2^K$ .
- For **balanced** designs, can express as  $\hat{\beta}_j = \frac{1}{2^{K-1}} \sum_i (y_{i_j+} y_{i_j-}) = \frac{1}{2^{K-1}} \sum_i \delta_{i_j+:i_j-}$ .

	<b>x</b> .1	<b>X</b> .2	• • •	$\mathbf{x}_{.j}$	• • •	$\mathbf{x}_{.K}$	У
$\mathbf{x}_{1.}$				• • •		• • •	
$\mathbf{x}_{i_j+.}$	<i>X</i> <sub><i>i</i><sub>j</sub> 1</sub>	$x_{i_j2}$		+1	• • •	$x_{i_jK}$	$y_{i_j}+$
• • •		• • •	• • •		• • •		
$\mathbf{x}_{i_j}$ .	<i>X</i> <sub><i>i</i><sub>j</sub> 1</sub>	$x_{i_j2}$		-1	• • •	$x_{i_jK}$	Уij—
• • •					• • •		
$\mathbf{X}_{2}^{K}$							



- For **orthogonal** designs  $\hat{\beta} = (X'X)^{-1}X'$ **y** becomes  $\hat{\beta}_j = \mathbf{x}'_{.j}\mathbf{y}/2^K$ .
- For **balanced** designs, can express as  $\hat{\beta}_j = \frac{1}{2^{K-1}} \sum_i (y_{i_j+} y_{i_j-}) = \frac{1}{2^{K-1}} \sum_i \delta_{i_j+;i_j-}$ .
- ullet Key observations: Only involves **simulation output**  $oldsymbol{\delta}$ ; **other parameters** *ceteris paribus*.

	<b>x</b> .1	<b>X</b> .2	• • •	$\mathbf{x}_{.j}$	• • •	$\mathbf{x}_{.K}$	У
$\mathbf{x}_{1.}$				• • •		• • •	
$\mathbf{x}_{i_j+.}$	<i>X</i> <sub><i>i</i><sub>j</sub> 1</sub>	$x_{i_j2}$		+1	• • •	$x_{i_jK}$	$y_{i_j}+$
• • •		• • •	• • •		• • •		
$\mathbf{x}_{i_j}$ .	<i>X</i> <sub><i>i</i><sub>j</sub> 1</sub>	$x_{i_j2}$		-1	• • •	$x_{i_jK}$	Уij—
• • •					• • •		
$\mathbf{X}_{2}^{K}$							



- For **orthogonal** designs  $\hat{\boldsymbol{\beta}} = (X'X)^{-1}X'\mathbf{y}$  becomes  $\hat{\beta}_j = \mathbf{x}'_{.j}\mathbf{y}/2^K$ .
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- ullet Key observations: Only involves **simulation output**  $oldsymbol{\delta}$ ; **other parameters** *ceteris paribus*.
- For Full Factorial design **simply replace**  $\delta_{i_j+;i_j-}$  with  $\delta_{i_j+;i_j-}^F = \sum_{k=1}^K w_{i_j+;i_j-;k} |\sigma_{i_j+}(k) \sigma_{i_j-}(k)|$ .

	<b>x</b> .1	<b>X</b> .2	• • •	$\mathbf{x}_{.j}$		$\mathbf{x}_{.K-p}$	$\mathbf{x}_{.K-p+1}$		$\mathbf{x}_{.K}$	У
<b>x</b> <sub>1.</sub>						I				
$\mathbf{x}_{i_j+.}$	<i>x</i> <sub>ij</sub> 1	$x_{i_j2}$		+1	• • •	$x_{i_jK-p}$	$X_{i_j+K-p+1}$		$x_{i_j+K}$	$y_{i_j+}$
		• • •	• • •	• • •	• • •			• • •		
$\mathbf{x}_{i_j}$ .	<i>X<sub>ij</sub></i> 1	$x_{i_j2}$	• • •		• • •	$x_{i_jK-p}$	$X_{ij}-K-p+1$	• • •	$x_{i_j-K}$	Уij—
• • • •	• • •	• • •		• • •	• • •	• • •		• • •	• • •	
$\mathbf{x}_{2}^{\kappa-\rho}$ .				• • •	• • •		•••	• • •	• • • •	

						_				
	<b>x</b> .1	<b>X</b> .2		$\mathbf{x}_{.j}$		$\mathbf{x}_{.K-p}$	$\mathbf{x}_{.K-p+1}$	• • •	$\mathbf{x}_{.K}$	у
<b>x</b> <sub>1.</sub>		• • •								
			• • •	• • •						
$\mathbf{x}_{i_j+.}$	<i>x</i> <sub>ij</sub> 1	$x_{i_j2}$	• • •	+1	• • •	$x_{i_jK-p}$	$x_{i_j+K-p+1}$	• • •	$x_{i_j+K}$	$y_{i_j+}$
			• • •		• • •			• • •		
$\mathbf{x}_{i_j}$	<i>X<sub>ij</sub></i> 1	$x_{i_j2}$	• • •	-1	• • •	$x_{i_jK-p}$	$X_{i_j-K-p+1}$	• • •	$x_{i_j-K}$	y <sub>ij</sub> _
		• • •				• • • •			• • •	
$\mathbf{x}_{2}^{\kappa-\rho}$ .					• • •		•••		• • • •	

• For orthogonal and balanced designs, **still** 
$$\hat{\beta}_j = \frac{1}{2^{K-\rho-1}} \sum_i (y_{i_j+} - y_{i_j-}) = \frac{1}{2^{K-\rho-1}} \sum_i \delta_{i_j+;i_j-}$$
.

						_				
	<b>x</b> .1	<b>X</b> .2		$\mathbf{x}_{.j}$		$\mathbf{x}_{.K-p}$	$\mathbf{x}_{.K-p+1}$	• • •	$\mathbf{x}_{.K}$	У
<b>x</b> <sub>1.</sub>		• • •								
			• • •	• • •						
$\mathbf{x}_{i_j+.}$	<i>X</i> <sub><i>i</i><sub>j</sub> 1</sub>	$x_{i_j2}$	• • •		• • •	$x_{i_jK-p}$	$X_{i_j+K-p+1}$	• • •	$x_{i_j+K}$	$y_{i_j}+$
			• • •		• • •			• • •		
$\mathbf{x}_{i_j}$	<i>X</i> <sub><i>i</i><sub>j</sub> 1</sub>	$x_{i_j2}$	• • •	-1	• • •	$x_{i_jK-p}$	$x_{i_j-K-p+1}$	• • •	$x_{i_j-K}$	y <sub>ij</sub> _
			• • •		• • •					
$\mathbf{x}_{2}^{\kappa-\rho}$ .					• • •		•••		• • • •	

- For orthogonal and balanced designs, **still**  $\hat{\beta}_j = \frac{1}{2^{K-p-1}} \sum_i (y_{i,+} y_{i,-}) = \frac{1}{2^{K-p-1}} \sum_i \delta_{i,+;i,-}$ .
- But now **careful enumeration** of pairs of rows for  $\delta_{i_1+i_2}^F$  for each parameter  $j=1,\ldots,K$ .

				_		_				
	<b>x</b> .1	<b>X</b> .2	• • •	$\mathbf{x}_{.j}$	• • •	$\mathbf{x}_{.K-p}$	$\mathbf{x}_{.K-p+1}$		$\mathbf{x}_{.K}$	У
<b>x</b> <sub>1.</sub>									• • •	
		• • •	• • •						• • •	
$\mathbf{x}_{i_j+.}$	<i>X</i> <sub><i>i</i><sub>j</sub>1</sub>	$x_{i_j2}$	• • •	+1	• • •	$x_{i_jK-p}$	$x_{i_j+K-p+1}$	• • •	$x_{i_j+K}$	$y_{i_j+}$
			• • •		• • •			• • •	• • •	
$\mathbf{x}_{i_j}$ .	<i>X</i> <sub><i>i</i><sub>j</sub> 1</sub>	$x_{i_j2}$	• • •	-1	• • •	$x_{i_jK-p}$	$x_{i_j-K-p+1}$	• • •	$x_{i_j-K}$	<i>y<sub>ij</sub></i> —
$\mathbf{x}_{2}^{\kappa-p}$ .		• • •			• • •			• • •	• • • •	

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- But now **careful enumeration** of pairs of rows for  $\delta_{i,+:i,-}^F$  for each parameter  $j=1,\ldots,K$ .
- For first K p parameters, **simply ignore** remaining p columns and  $i_i + i_j c$  hosen as before.

	<b>x</b> .1	<b>X</b> .2		$\mathbf{x}_{.j}$		$\mathbf{x}_{.K-p}$	$\mathbf{x}_{.K-p+1}$	• • •	$\mathbf{x}_{.K}$	У
<b>x</b> <sub>1.</sub>		• • •								
			• • •	• • •						
$\mathbf{x}_{i_j+.}$	<i>X</i> <sub><i>i</i><sub>j</sub> 1</sub>	$x_{i_j2}$	• • •		• • •	$x_{i_jK-p}$	$X_{i_j+K-p+1}$	• • •	$x_{i_j+K}$	$y_{i_j}+$
			• • •		• • •			• • •		
$\mathbf{x}_{i_j}$	<i>X</i> <sub><i>i</i><sub>j</sub> 1</sub>	$x_{i_j2}$	• • •	-1	• • •	$x_{i_jK-p}$	$x_{i_j-K-p+1}$	• • •	$x_{i_j-K}$	y <sub>ij</sub> _
			• • •		• • •					
$\mathbf{x}_{2}^{\kappa-\rho}$ .					• • •		•••		• • • •	

- For orthogonal and balanced designs, **still**  $\hat{\beta}_i = \frac{1}{2K-p-1} \sum_i (y_{i,+} y_{i,-}) = \frac{1}{2K-p-1} \sum_i \delta_{i,+;i,-}$ .
- But now **careful enumeration** of pairs of rows for  $\delta_{h+j}^F$  for each parameter  $j=1,\ldots,K$ .
- For first K p parameters, **simply ignore** remaining p columns and  $i_i + i_j$  chosen as before.
- For K p + j-th parameter, use its column for one of the first K p columns whose parameter was used to construct it (via the generator words).

#### **Summary**

### **Summary**

Sensitivity Analysis via low-order polynomial meta-models fit using OLS regression and Factorial designs:

- Combat simulations often depart from (all) standard NIID assumptions:
  - Non-independently distributed (via use of CRN).
  - Non-identically distributed (variance depends on design point).
  - Non-normally distributed (skewness).
- Kleijnen (2015) text suggested different remedies, but they are actually equivalent.
- Kleijnen (2015) also incorrectly derived lack-of-fit F-statistics in white-noise case.

# **Summary**

Sensitivity Analysis via low-order polynomial meta-models fit using OLS regression and Factorial designs:

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- Kleijnen (2015) text suggested different remedies, but they are actually equivalent.
- Kleijnen (2015) also incorrectly derived lack-of-fit F-statistics in white-noise case.

Sensitivity Analysis of Analysis of Alternatives new research topic:

- Borrow distance metrics from Information Retrieval algorithm comparisons.
- ullet Exploit  $\delta$ -structure and *ceteris paribus* principle of Full Factorial designs.
- Future work: test effectiveness of approach; examine other orthogonal/balanced designs; improve ranking sensitivity measure.